# Defining single extreme weather events in a climate perspective 

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## Context

The analysis of single extreme weather events relates to:

- climate monitoring;
- physical understanding;
- estimation of return periods;
- attribution to climate change.

For the latter, one compares the probability of the event occurring:

- in the factual world ( $p_{1}$ );
- in a counter-factual world, e.g. non-anthropized ( $p_{0}$ ).

One defines the Risk Ratio and the Fraction of Attributable Risk as:

$$
R R=\frac{p_{1}}{p_{0}} \quad \text { and } \quad F A R=\frac{p_{1}-p_{0}}{p_{1}}=1-\frac{1}{R R}
$$

## The four steps of event definition

1. Select the variable $(X)$.

Usually straightforward - not crucial here.
2. Define the class of events.

Here, traditional "risk-based" approach, i.e. events equally or more intense than the observed one: $\operatorname{Pr}\left\{X \geq x_{0}\right\}$ with $x_{0}$ the event value.
N.B. Alternative "storyline" approach: events of about the same intensity - not appropriate for probabilistic framework since: $\operatorname{Pr}\left\{x_{0}-\varepsilon \leq X \leq x_{0}+\varepsilon\right\} \xrightarrow{\varepsilon \rightarrow 0} 0$.
3. Define the level of conditioning.
$\operatorname{Pr}\left\{X \geq x_{0} \mid Y \in \Omega\right\}$ with $Y$ a concurrent climate variable (e.g. SST, atmospheric circulation, ENSO ) or the time of the year (e.g. winter heat wave)?

Here only the calendar conditioning is explored (relevant for climate monitoring).
4. Define the spatio-temporal scale.

The main topic of this talk.

## Why spatio-temporal scale matters

Example of the European heat-wave of summer 2003 (EHW03):

- Stott et al. (2004): EHW03 becomes a cold extreme after 2050.
- Beniston (2007): EHW03 remains a hot event in 2100.

The difference? Seasonal/European vs. daily/local temperature anomalies.
a) JJA T Europe


Stott et al., Nature, 2004 (SRES A2 scenario).
b) Daily $T$ Basel


## Why spatio-temporal scale matters

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## Choice of spatio-temporal scale

Most of the time: arbitrary.
Authors use predefined areas (e.g. a local station, a national territory) and periods (e.g., a day, a month, a season), and/or their own expertise.

Problem \#1: this may not faithfully portray the event / be biased by our perception.
Problem \#2: different definitions of the "same" event may lead to different attribution statements (see EHW03 example).

Our idea: select the scale at which the event has been the most extreme, i.e. minimize the factual probability:

$$
p_{1}=\operatorname{Pr}\left\{X^{\left(t_{1}\right)} \geq x_{t_{1}}\right\}
$$

with $X^{\left(t_{1}\right)}$ the random variable describing the temperature distribution at time $t_{1}=2003$, and $x_{t_{1}}$ the observed 2003 value.

## Optimizing the time window

Example: Daily T at Paris-Montsouris station for Jun-Jul-Aug 2003.


Data: Météo-France.

Question: Over which time window is the anomaly the most extreme?

1. Aug 11 (1 day);
2. Aug 5-12 (1 week);
3. Aug 2-17 (2 weeks);
4. August (1 month);
5. June (1 month);
6. Jun-Jul-Aug (1 season).

## Optimizing the time window - Calendar method

For each time window $\llbracket d_{1}, d_{2} \rrbracket$ :

- we consider the observed time series $x_{t}$ \& the climate change $x_{t}^{*}$ at this location;

N.B. $x_{t}^{*}=$ smoothed multi-model mean of CMIP5 JJA temperatures (common to all time windows but location-dependent).


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For each time window $\llbracket d_{1}, d_{2} \rrbracket$ :

- we consider the observed time series $x_{t}$ \& the climate change $x_{t}^{*}$ at this location; - we correct for climate change before and after $t_{1}=2003: x_{t}^{\left(t_{1}\right)}=x_{t}-\left(x_{t}^{*}-x_{t_{1}}^{*}\right)$;


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- we estimate $p_{1}$ from $x_{t}^{\left(t_{1}\right)}$, assuming $X^{\left(t_{1}\right)}$ follows a Gaussian distribution;

T Paris for Aug 1 - Aug 10



T Paris pdf for Aug 1 - Aug 10

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- we estimate $p_{1}$ from $x_{t}^{\left(t_{1}\right)}$, assuming $X^{\left(t_{1}\right)}$ follows a Gaussian distribution;
- we also estimate $p_{0}$ by correcting wrt. $t_{0}=1950$ (our counter-factual world).

T Paris for Aug 1 - Aug 10


T Paris pdffor Aug 1 - Aug 10

N.B. $x_{t}^{*}=$ smoothed multi-model mean of CMIP5 JJA temperatures (common to all time windows but location-dependent).

## Optimizing the time window - Result

The most extreme anomaly is found for Aug 5-12 ( $p_{1}=4 \times 10^{-6}, 250000 \mathrm{y}$ ).


## Calendar vs. annual maxima

The calendar approach is relevant for climate monitoring (seasonal context), but the obtained $p_{1}$ should not be interpreted as a formal return period.

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The calendar approach is relevant for climate monitoring (seasonal context), but the obtained $p_{1}$ should not be interpreted as a formal return period. Alternative approach: consider $x_{t}^{\left(t_{1}\right)}$ as the time series of annual maxima, (now assuming that $X^{\left(t_{1}\right)}$ follows a Gumbel distribution).

T Paris pdf for Aug 1 - Aug 10

N.B. The question becomes: Over which time window is the temperature the most extreme?

## Calendar vs. annual maxima - Result

The most extreme temperature is found for Aug 4-12 ( $p_{1}=0.008,125 \mathrm{y}$ ). Hot anomalies distant from the annual cycle peak disappear (e.g. June).



## A compromise: local maxima

Idea : limit the search of annual maxima to a calendar neighborhood, i.e. consider $x_{t}^{\left(t_{1}\right)}$ as the time series of local maxima.
Here we use $\pm 7$ days; similar to what is done for establishing record values.



## Optimizing the space window

Idea (simple): repeat the procedure for an ensemble of spatial domains...
Here: squared or near-squared domains including Paris / included in Europe. Observations: E-OBS interpolated onto a $2.5 \times 2.5^{\circ}$ grid.


Alternative methods: successive grouping of countries or hierarchical collection of regions proposed by D.A. Stone (Climatic Change, submitted).

## Optimizing the space window - Result

Annual-maxima / local-maxima: minimum $p_{1}(0.005,200 \mathrm{y})$ is found for Aug 2-13 over France \& Spain (12 days, $7 \times 5$ domain).
Calendar approach: other minimum at smaller scale ( 8 days, $2 \times 1$ domain).


Domain size in number of grid points


$x$-axis: size of the space window from local (Paris) to the entire Europe. $y$-axis: size of the time window from 1 day to the entire season ( 92 days).

## Does it bias the FAR?

$$
p_{1}=\operatorname{Pr}\left\{X^{\left(t_{1}\right)} \geq x_{t_{1}}\right\}, p_{0}=\operatorname{Pr}\left\{X^{\left(t_{0}\right)} \geq x_{t_{1}}\right\} \text { and FAR }=1-p_{0} / p_{1} .
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No. For this event, the FAR increases with spatio-temporal scale. It is maximum for the scale chosen by Stott et al. (2004).




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The FAR responds to the signal-to-noise ratio of the human-induced change. For temperatures, the signal (warming) is rather uniform across scales, while the noise (variability) is stronger for small spatio-temporal scales.

## Another temperature event

The European heat-wave of summer 2018.
Result: Sweden-Finland, Jul 14 - Aug 2, estimated return period 50 y.

- Higher $p_{1}$ than 2003: less extreme event.
. Higher FAR values: stronger signal-to-noise ratio in 2018 vs. 2003.
. Same behavior for the FAR: it increases with spatio-temporal scale.


Yiou et al., BAMS report on 2018 extremes, in prep.

## A precipitation event

The intense rainfall in Boulder, Colorado, September 2013.
Method: annual maxima, with a different correction for climate change.

1. We estimate the local long-term $T$ change $x_{t}^{*}$ (in K, CMIP5).
2. We estimate the scaling of the annual n-day $P$ maxima (in \% par K, CMIP5).
$\longrightarrow 2.5 \% / K$ for 1-day maxima, $0.7 \% / K$ for 92-day maxima.
3. We rescale the $P$ annual max time series wrt. $2013\left(p_{1}\right)$ or $1901\left(p_{0}\right)$.
4. We use GEV distributions with shape parameter $\xi=0.1$ across all time windows.


Data: GHCN daily data at Boulder station + regridded at $0.1 \times 0.1^{\circ}$ by M . Hoerling.

## A precipitation event - Result

- $p_{1}$ is found to be minimum for Boulder local station, Sep 11-15.
. Large estimated return period: $p_{1}=7 \times 10-5$, i.e. 15000 y .
- Rather small FAR values, typically between 10 and $25 \%$.
$\hookrightarrow$ Consistent with previous attribution studies (Hoerling et al., 2014; Eden et al. 2016).
For this event, the FAR decreases with spatio-temporal scale.
The signal-to-noise ratio is more complex for $P$ than for $T$.



## Summary

Select the space-time window that maximizes the event rarity (minimizes $p_{1}$ ) provides an as-objective-as-possible event definition.

Maximizing the rarity does not systematically maximize (or minimize) the attributable risk, contrarily to some arbitrary definitions.

Using $p_{1}$ allows to compare the rarity of different events and/or select the events that have been the most extreme within a year (e.g. for BAMS reports).

We have used very simple detrending + probability estimation procedures, future work may involve including more sophisticated techniques.

Cattiaux, J. and A. Ribes, Defining single extreme weather events in a climate perspective, Bulletin of the American Meteorological Society, 99, 1557-1568. doi:10.1175/BAMS-D-17-0281.1

